

Heteroskedasticity

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Abstract

We can correct for heteroskedasticity even when estimating a mean.

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1. Introduction asdfadfds

Serial correlation is about having a random sample. Heteroskedasticity is about having data of different quality.

2. The Weighted Mean

We want to estimate the mean μ from N observations $y_i = \mu + \epsilon_i$, where the ϵ_i are independently distributed with mean 0 and variance σ_i^2 . Suppose we estimate the population mean by using a weighted mean:

$$\hat{\mu}^* \equiv \frac{\sum_{i=1}^N w_i y_i}{N}, \quad (1)$$

where we choose w_i so that $\sum_{i=1}^N w_i = N$. We could use $w_i = 1$, for example, which is the same as the unweighted mean.

This is an unbiased estimator.

$$E\hat{\mu}^* = E \frac{\sum_{i=1}^N w_i y_i}{N} = E \frac{\sum_{i=1}^N (w_i \mu + w_i \epsilon_i)}{N} = \frac{\mu \sum_{i=1}^N w_i}{N} + E \frac{\sum_{i=1}^N w_i \epsilon_i}{N} = \frac{\mu}{N} N + 0 = \mu \quad (2)$$

The variance of the estimator is

$$\begin{aligned} \text{Var}(\hat{\mu}^*) &= E \frac{\left(\frac{\sum_{i=1}^N w_i y_i}{N} - \mu \right)^2}{N} \\ &= E \frac{\left(\frac{\sum_{i=1}^N w_i \mu + w_i \epsilon_i}{N} - \mu \right)^2}{N} \\ &= E \frac{\left(\mu + \frac{\sum_{i=1}^N w_i \epsilon_i}{N} - \mu \right)^2}{N} \\ &= \frac{\sum_{i=1}^N w_i^2 \sigma_i^2}{N} \\ &= \frac{\sum_{i=1}^N w_i^2 \sigma_i^2}{N} \end{aligned} \quad (3)$$

Notice, in particular, that if $w_i = 1$ for all i , then

$$\text{Var}(\hat{\mu}^*) = \frac{\sum_{i=1}^N \sigma_i^2}{N} = \frac{\sigma^2}{N}. \quad (4)$$

But we can do better than that. Let's choose w_i to minimize the variance. How do we do that? With a giant lagrangian? We want to minimize the variance such that $\sum w_i = N$. So we solve

$$\text{Minimize}_{w_i} \frac{\sum_{i=1}^N w_i^2 \sigma_i^2}{N^2} - \lambda (\sum w_i - N) \quad (5)$$

This has the first order conditions for each i ,

$$\frac{2w_i \sigma_i^2}{N^2} - \lambda = 0, \quad (6)$$

so

$$w_i = \frac{N^2 \lambda}{2\sigma_i^2} \quad (7)$$

or

$$\lambda = \frac{2\sigma_i^2 w_i}{N^2} = \frac{2\sigma_j^2 w_j}{N^2} \quad (8)$$

Thus, $\sigma_i^2 w_i = \sigma_j^2 w_j$ and $\frac{w_i}{w_j} = \frac{\sigma_j^2}{\sigma_i^2}$. A weight that satisfied that first order condition and the constraint is:

$$w_i^* = \frac{\bar{\sigma}^2}{\sigma_i^2} \quad (9)$$

This puts more weight on the better data, the observations with lower variance.

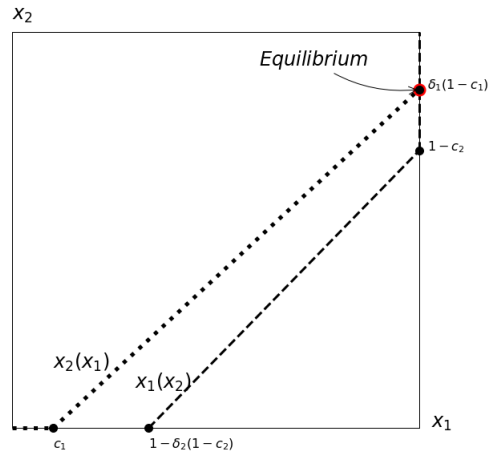
The average variance of the residuals, the sample variance, is (with the N-1 correction so as to be unbiased):

$$\begin{aligned} \text{Var}(\hat{\epsilon}_i^*) &= E \frac{\sum_{i=1}^N \left(\frac{\sum_{j=1}^N w_j y_j}{N} - y_i \right)^2}{N-1} \\ &= E \frac{\sum_{i=1}^N \left(\frac{\sum_{j=1}^N (w_j \mu + w_j \epsilon_j)}{N} - \mu - \epsilon_i \right)^2}{N-1} \\ &= E \frac{\sum_{i=1}^N \left(\frac{\sum_{j=1}^N w_j \epsilon_j}{N} - \epsilon_i \right)^2}{N-1} \\ &= E \sum_{i=1}^N \left(\frac{(\sum_{j=1}^N w_j \epsilon_j)^2}{N^2(N-1)} - 2 \frac{\sum_{j=1}^N w_j \epsilon_j \epsilon_i}{N(N-1)} + \frac{\epsilon_i^2}{N-1} \right) \\ &= \sum_{i=1}^N \left(\frac{\sum_{j=1}^N w_j^2 \sigma_j^2}{N^2(N-1)} - \frac{2w_i \sigma_i^2}{N(N-1)} + \frac{\sigma_i^2}{N-1} \right) \\ &= \sum_{i=1}^N \left(\frac{\sum_{j=1}^N w_j^2 \sigma_j^2}{N^2(N-1)} - \frac{2N w_i \sigma_i^2}{N^2(N-1)} + \frac{N^2 \sigma_i^2}{N^2(N-1)} \right) \\ &= \frac{N^2 \bar{\sigma}^2}{N^2(N-1)} - \frac{2N^2 \bar{\sigma}^2}{N^2(N-1)} + \frac{N^3 \bar{\sigma}^2}{N^2(N-1)} \text{sd}f \text{sd}f \text{sd}d \\ &= -\frac{\bar{\sigma}^2}{N-1} + \frac{N \bar{\sigma}^2}{N-1} \text{sd}f \text{sd}f d \\ &= \frac{(N-1) \bar{\sigma}^2}{N-1} \text{sd}f \text{sd}f \\ &= \bar{\sigma}^2 \text{sd}f d \end{aligned} \quad (10)$$

How do we estimate σ_i^2 ?

4. Template stuff

FIGURE 1:
SDFSAFSDFAFD REACTION CURVES x_1 AND x_2



References

Kennedy

White